

# Averaging Mortality Rate Forecasts Across Fitting Periods

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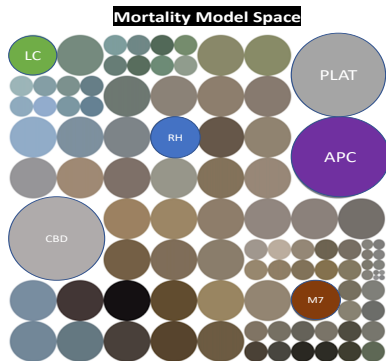
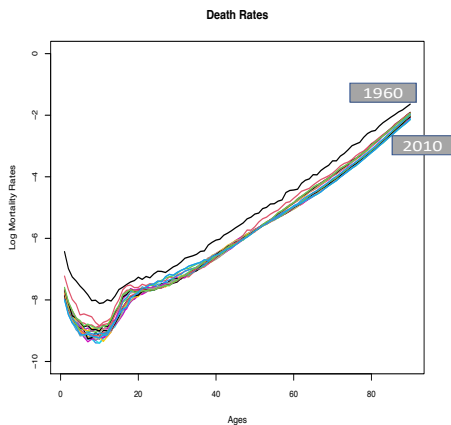
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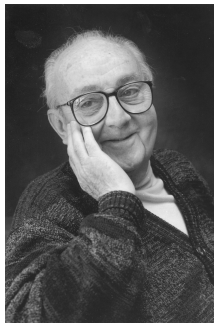
# Model Selection Dilemma



- What mortality model is likely to perform best?

# Model Combination: Motivation

- George Edward Box is one of the great statistical minds of the 20th century.
- George Edward Box said **all models are wrong, but some are useful.**



# Model combination is an alternative to model selection



- Bayesian Model Averaging (Barigou et al. 2021).
- Model Confidence Set (Shang and Haberman 2018).
- Stacked regression ensemble (Wolpert 1992).

**Model combination formulation:**

$$\ln(\widehat{\mu}(x, t + h))_{\text{comb}} = \sum_{m=1}^M w_m \ln(\widehat{\mu}_m(x, t + h)).$$





## CoMoMo Package

- Use CoMoMo to estimate combination weights using stacked regression ensemble, Bayesian Model Averaging (Kessy et al. 2022).
- Access our paper



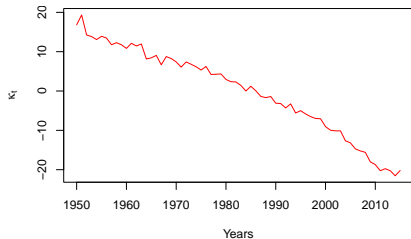
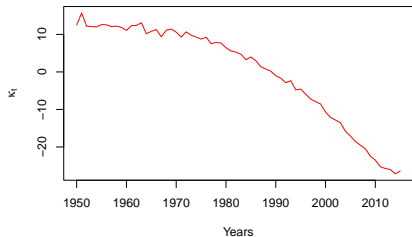
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# Fitting Period Selection Risk

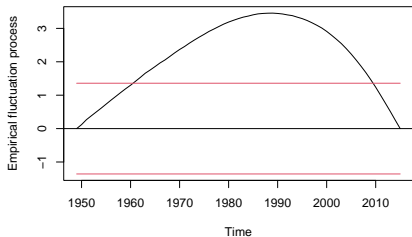
- The accuracy of mortality rate forecasts depends on the length of the fitting period used for estimating the model parameters (Booth, Maindonald, and Smith 2002).
- Choosing the wrong fitting period leads to bad decision-making.
- For example, when we underestimate the life expectancy by one year, the pension liabilities can increase by up to 5% (Bergeron-Boucher and Kjærsgaard 2021).

# Fitting Period Selection Dilemma

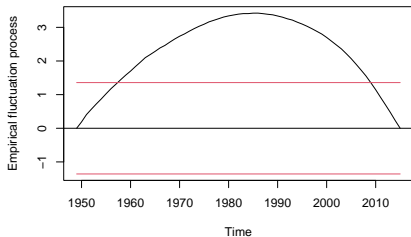
# Structural Changes in Mortality Patterns



**OLS-based CUSUM test**



**OLS-based CUSUM test**



# Existing Methods of Selecting Optimal Fitting Period

# We Ask

- Do the linear mortality models generate statistically different out-of-sample mean squared errors based on multiple fitting periods selected using linear regularization methods?
- Do the mortality rate forecasts averaged from the same linear mortality model calibrated over different fitting periods selected using linear regularization methods have lower out-of-sample mean squared errors than those based on the “longest” calibration period?

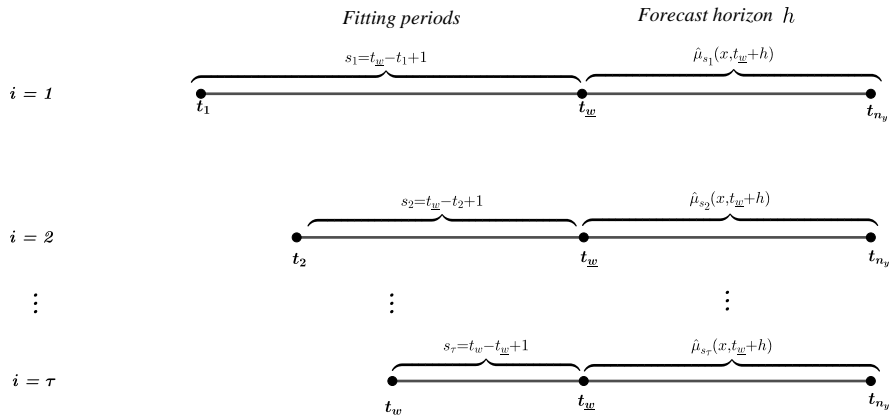


# Averaging Across Fitting Periods

- Presence of structural breaks and mortality time index is non-linear.



# Combine Forecasts Across Fitting Periods



# Averaging Across Fitting Periods

- Fit a mortality model using the fitting periods  $\{t_1 : t_{\underline{w}}, t_2 : t_{\underline{w}}, \dots, t_{\omega} : t_{\underline{w}}\}$ .
- Define the sample size  $s_i$  as  $s_i = t_{\underline{w}} - t_i + 1, \forall i = 1, \dots, \tau$  and  $\tau = |\{t_1 : t_{\underline{w}}, t_2 : t_{\underline{w}}, \dots, t_{\omega} : t_{\underline{w}}\}|$ .
- Use the fitted mortality model to generate the future mortality rates  $\hat{\mu}_{s_1}(x, t_{\underline{w}} + h), \dots, \hat{\mu}_{s_{\tau}}(x, t_{\underline{w}} + h)$  for age  $x \in [x_1, x_{n_a}]$  at time  $t_{\underline{w}} + h$ , where  $h$  is the forecasting horizon  $\forall h = 1, \dots, H$ .
- Model combination formulation (Pesaran and Timmermann 2007):

$$\ln(\hat{\mu}(x, t_{\underline{w}} + h))_{\text{comb}} = \sum_{i=1}^{\tau} w_{s_i}(h) \ln \hat{\mu}_{s_i}(x, t_{\underline{w}} + h), \forall h = 1, \dots, H.$$

# Simple Fitting Periods Averaging

- Equal weighting multiple fitting periods is empirically robust to structural breaks of unknown break dates and sizes (Pesaran and Timmermann 2007).
- Equally weighted combined mortality rate forecast  $\ln(\hat{\mu}(x, t_{\underline{w}} + h))_{\text{SFPA}}$  for different forecasting horizons  $h$  from  $\tau$  multiple starting points is defined by the simple forecast combination rule

$$\ln(\hat{\mu}(x, t_{\underline{w}} + h))_{\text{SFPA}} = \sum_{i=1}^{\tau} \frac{1}{\tau} \ln \hat{\mu}_{s_i}(x, t_{\underline{w}} + h), \forall h = 1, \dots, H.$$

## Location Based Weights

- Recent mortality data contribute more to forecasting future mortality rates because they reflect the most recent dynamics of mortality evolution (Hyndman and Shang 2009).
- The combination weight should be proportional to the location  $i \in \{t_1, \dots, t_w\}$  such that we assign heavier weights to the out-of-sample forecasts based on the recent mortality data.
- For the forecasting horizon  $h$ ,  $\forall h = 1, \dots, H$ ,

$$\ln(\hat{\mu}(x, t_w + h))_{\text{LOC}} = \sum_{i=1}^{\tau} \frac{i}{\sum_{i=1}^{\tau} i} \ln \hat{\mu}_{s_i}(x, t_w + h)$$

- Exponentially decaying weights (Hyndman and Shang 2009)

$$\ln(\hat{\mu}(x, t_w + h))_{\text{EXP}} = \sum_{i=1}^{\tau} \lambda(1 - \lambda)^{s_i} \ln \hat{\mu}_{s_i}(x, t_w + h), \forall h = 1, \dots, H.$$

# Averaging across Windows and Models

- To account for model uncertainty, multiple structural changes, and the impact of the fitting period.
- Let the  $h$ -year-ahead mortality rate forecasts from mortality models  $m = 1, \dots, M$  be  $\hat{\mu}_{s_i}^1(x, t_{\underline{w}} + h), \dots, \hat{\mu}_{s_i}^M(x, t_{\underline{w}} + h)$  estimated over the sample window  $s_i, \forall i = 1, \dots, \tau$  for age  $x \in [x_1, x_{n_a}]$  at time  $t_{\underline{w}} + h$ .
- Combine mortality rate forecasts from across mortality models and different fitting periods  $\ln(\hat{\mu}(x, t_{n_y} + h))_{\text{AveAve}}$  using

$$\ln(\hat{\mu}(x, t_{\underline{w}} + h))_{\text{AveAve}} = \frac{1}{M} \sum_{m=1}^M \sum_{i=1}^{\tau} w_{s_i}(h) \ln \hat{\mu}_{s_i}^m(x, t_{\underline{w}} + h), \forall h = 1, \dots, H$$

# Generate Metadata Using Cross-Validation

- Generate cross-validation predictions using the block-cross validation (SriDaran et al. 2021).
- Produce the cross-validation predictions from multiple starting years,  $\{t_1 : t_w, t_2 : t_w, \dots, t_w : t_w\}$  using a given mortality model.
- Train the mortality model on the window  $t_1 : t_w^*$  to generate cross-validation predictions,  $\mathbf{z}_{h,s_1} \forall h = 1, 2, \dots, H$ .
- Train the mortality model on the window  $t_2 : t_w^* + 1$  to generate cross-validation predictions,  $\mathbf{z}_{h,s_2} \forall h = 1, 2, \dots, H$ .
- Repeat this procedure until the calibration window is long enough to generate cross-validation predictions for the longest horizon which is  $H = 15$ .

# Generate Metadata Using Cross-Validation

- When  $h = 1, \dots, H$ , let  $\mathbf{z}_{h=h, s_i}$  for  $i = 1, 2, \dots, \tau$  forms the matrix

$$\mathcal{M}_{h=1} = [\mathbf{z}_{h=1, s_1}, \mathbf{z}_{h=1, s_2}, \dots, \mathbf{z}_{h=1, s_\tau}].$$

- Then for  $h = 1, \dots, H$ , stack  $\mathcal{M}_{h=1}, \mathcal{M}_{h=2}, \dots, \mathcal{M}_{h=H}$  in row order on top of each other to form a new matrix  $\mathcal{M}^+$  given as

$$\mathcal{M}^+ = \begin{bmatrix} \mathcal{M}_1 \\ \vdots \\ \mathcal{M}_H \end{bmatrix}.$$

- The matrix  $\mathcal{M}^+$  along with the observed mortality rates  $\mathbf{y}$  forms the metadata, and we need to solve the following linear regression

$$\mathbf{y} = \mathcal{M}^+ \mathbf{w},$$

where  $\mathbf{w} = [w_1, \dots, w_\tau]'$  and we estimate  $\mathbf{w}$  using regularized linear regression methods.



# Metadatda

	$\hat{\mu}_{1950}$	$\hat{\mu}_{1951}$	$\hat{\mu}_{1952}$	$\hat{\mu}_{1953}$	...	$\hat{\mu}_{1970}$	Actual
1	-4.91	-4.90	-4.87	-4.75	...	-4.91	-4.93
2	-4.87	-4.91	-4.85	-4.73	...	-4.90	-4.93
3	-4.86	-4.92	-4.85	-4.73	...	-4.90	-4.89
$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$
1197	-1.48	-1.50	-1.49	-1.38	...	-1.48	-1.49
1198	-1.50	-1.46	-1.46	-1.38	...	-1.44	-1.44
1199	-1.50	-1.46	-1.47	-1.36	...	-1.42	-1.52

$$\Rightarrow Y = Z\mathbf{w}$$

# Regularization Methods to Estimate Weights

- Use regularization methods such as non-negative least regression to handle correlated forecasts in metatadata (Breiman 2004).
- Based on Equation (20), determine the weighting coefficients  $\widehat{w}_{S_i} \quad \forall i = 1, 2, \dots, \tau$  as the minimizers of

$$\widehat{\mathbf{w}}^* = \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{j=1}^{M^*} \left( y_j - \sum_{i=1}^{\tau} w_{S_i} z_{S_i} \right)^2.$$

- In lasso regression, we estimate the weights as

$$\widehat{\mathbf{w}}^* = \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{j=1}^{M^*} \left( y_j - \sum_{i=1}^{\tau} w_{S_i} z_{S_i} \right)^2 + \lambda \sum_{m=1}^{\tau} |w_{S_i}|.$$

- Transform the weights estimated using the regularization methods such that the recent mortality data weigh more than the past data using

$$w_{S_i}^* = w_{S_i} (1 - w_{S_i})^{S_i}, \forall i = 1, \dots, \tau.$$

# Empirical Results

- Generalised Age-Period-Cohort mortality models (Villegas, Kaishev, and Millosovich 2018).

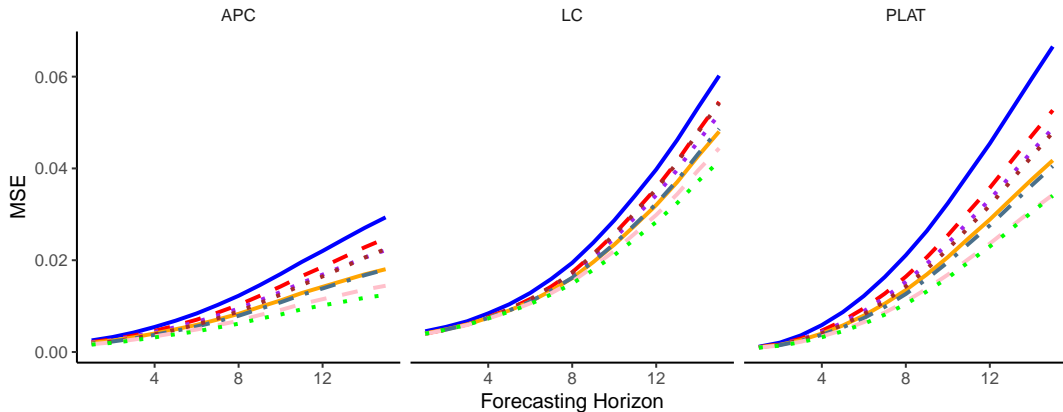
Model	Predictor ( $\eta_{xt}$ )	Parameters
LC	$\alpha_x + \beta_x^{(1)} \kappa_t^{(1)}$	$2n_a + n_y$
APC	$\alpha_x + \kappa_t^{(1)} + \gamma_c$	$n_a + n_y + n_b$
PLAT	$\alpha_x + \kappa_t^{(1)} + (\bar{x} - x) \kappa_t^{(2)} + (\bar{x} - x)^+ \kappa_t^{(3)} + \gamma_c$	$n_a + 3n_y + n_b$

Table 1: Generalized age-period-cohort mortality models.

- Human Mortality Database: England and Wales male mortality data (University of California Berkeley and the Max Planck Institute for Demographic Research 2020).
- Training set: 1950 to 1990, Test set: 1991 to 2015, and ages 50 – 89.

# Forecast Instability From Multiple Starting Periods

MSE of the one-step-ahead to 15-step-ahead mortality rate forecasts using different mortality models fitted to multiple fitting periods

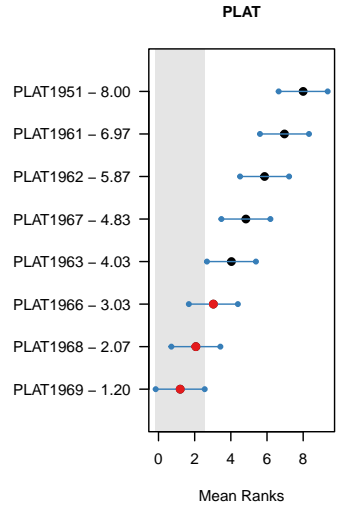
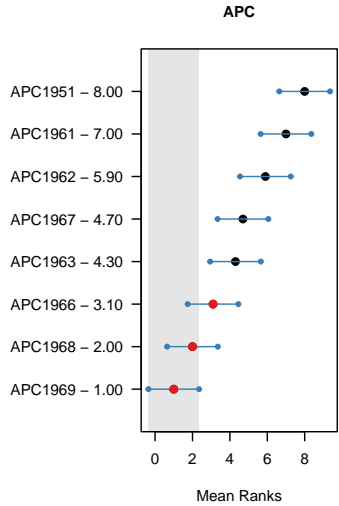
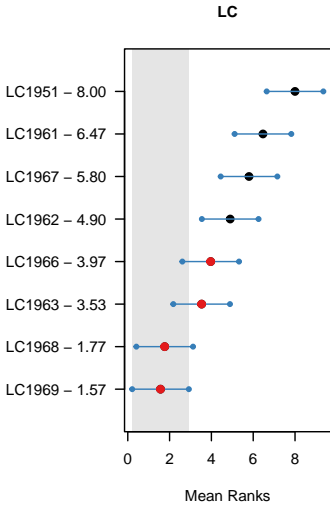


Multiple Starting Years

— 1951	· · · 1962	- - - 1966	- · - 1968
- · - 1961	— 1963	· · · 1967	· · · 1969

# Statistical Test of Forecast Errors

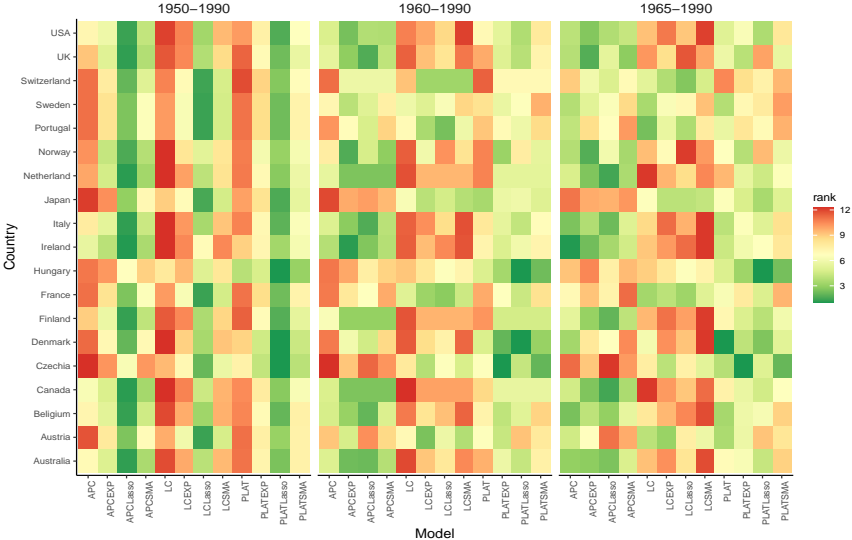
Multiple comparisons with the best style plot for the mortality models





# Average Ranks to 19 Populations

Average ranks of mortality models across different countries



## Conclusion and Future Research

- Mortality models fitted to different base periods generate out-of-sample mortality rate forecasts that differ statistically significantly.
- Empirical experiments on data from 19 male populations reveal averaging forecasts from different base periods further improves out-of-sample precision in the presence of structural breaks and when the mortality time index is non-linear.
- Future research could include the prediction of the occurrence of possible future structural changes while forecasting the mortality rates.
- Develop model combination approaches that generate prediction intervals from averaging multiple fitting periods.





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